

RESEARCH

Open Access



Aging and self-reported health in 114 Latin American cities: gender and socio-economic inequalities

Marianela Castillo-Riquelme^{1*}, Goro Yamada², Ana V. Diez Roux², Tania Alfaro¹, Sandra Flores-Alvarado¹, Tonatihu Barrientos³, Camila Teixeira Vaz⁴, Andrés Trotta⁵, Olga L. Sarmiento⁶ and Mariana Lazo^{2,7}

Abstract

Background: Understanding how urban environments influence people's health, especially as individuals age, can help identify ways to improve health in the rapidly urbanizing and rapidly aging populations.

Objectives: To investigate the association between age and self-reported health (SRH) in adults living in Latin-American cities and whether gender and city-level socioeconomic characteristics modify this association.

Methods: Cross-sectional analyses of 71,541 adults aged 25–97 years, from 114 cities in 6 countries (Argentina, Brazil, Colombia, Chile, El Salvador, and Guatemala), as part of the Salud Urbana en America Latina (SALURBAL) Project. We used individual-level age, gender, education, and self-reported health (SRH) data from harmonized health surveys. As proxies for socioeconomic environment we used a city-level socioeconomic index (SEI) calculated from census data, and gross domestic product (GDP) per-capita. Multilevel Poisson models with a robust variance were used to estimate relative risks (RR), with individuals nested in cities and binary SRH (poor SRH vs. good SRH) as the outcome. We examined effect modification by gender and city-level socioeconomic indicators.

Results: Overall, 31.4% of the sample reported poor SRH. After adjusting for individual-level education, men had a lower risk of poor SRH (RR = 0.76; CI 0.73–0.78) compared to women, and gender modified the association between age and poor SRH (p -value of interaction < 0.001). In gender stratified models, the association between older age and poor SRH was more pronounced in men than in women, and in those aged 25–65 than among those 65+ (RR/10 years = 1.38 vs. 1.10 for men, and RR/10 years = 1.29 vs. 1.02 for women). Living in cities with higher SEI or higher GDP per-capita was associated with a lower risk of poor SRH. GDP per-capita modified the association between age (25–65) and SRH in men and women, with SEI the interaction was less clear.

Conclusions: Across cities in Latin America, aging impact on health is significant among middle-aged adults, and among men. In both genders, cities with lower SEI or lower GDP per-capita were associated with poor SRH. More research is needed to better understand gender inequalities and how city socioeconomic environments, represented by different indicators, modify exposures and vulnerabilities associated with aging.

Keywords: Self-reported health, Aging, Latin-America, Multilevel analysis, Urban health, Gender, Inequalities

*Correspondence: marianelacastillo@hotmail.com; marianelacastillo@uchile.cl

¹ Doctoral Program in Public Health, School of Public Health, Faculty of Medicine, University of Chile, Avenida Independencia, 939 Santiago, Chile
Full list of author information is available at the end of the article

Background

By 2021, 56,6% of the global population lived in urban areas. Urbanization is especially high in Latin America, where the urban population is currently over 80%, together with North America among the highest across



regions of the world [1, 2]. In addition to urbanization, an important global demographic trend is aging, with Latin America and the Caribbean experiencing relatively rapid increases in older populations [3]. Between 1990 and 2019, the percentage of the population aged 65 years and older almost doubled in Latin America and the Caribbean (from 6 to 11%) [3]. By 2050, countries in Latin America and Asia are expected to experience the greatest absolute growth in the share of the population above 65 years [4]. Worldwide, many older adults live in urban areas. For example, in 2015 in the Organization for Economic Cooperation and Development (OECD) region, the share of the older population (65+) living in cities was 43% with trends showing more rapid growth in metropolitan areas compared to non-metropolitan areas [5, 6]. Hence, understanding how urban environments influence people's health across adulthood, and especially as individuals age, is an important research goal. Pursuing this goal can help identify ways to improve the health of older adults in the rapidly urbanizing and rapidly aging populations of Latin America and other similar regions.

Self-rated or self-reported health (SRH) has long been considered a reliable indicator of biological and mental health [7, 8], and a good predictor of morbidity and mortality [9, 10]. Mavaddat et al. for the UK context (study in 25,268 individuals aged 39–79) examined the strength of the association of single and multi-morbidity with SRH. They concluded that SRH provides a simple, integrative patient-centered assessment for evaluation of illness in the context of multiple chronic disease diagnoses [11].

SRH varies widely between [12–14] and within [15, 16] countries. A study that comprised 33 major metropolitan areas from eleven European countries reported that SRH was poorer in certain areas of the UK and Germany and better in areas of Sweden and Belgium [13]. Community and area socio-economic characteristics have also been linked to SRH within countries. For example, income inequality and poverty at the community or local area level, have been found to be associated with poor SRH in Chile [17], Colombia [18], United States [19] Wales [20] and Japan [21]. These findings suggest SRH can be influenced by place and living conditions.

Most studies on SRH differences have analyzed individuals within countries [22–24], with few studies conducted to understand how SRH varies across cities or is influenced by features of urban areas. The socioeconomic, physical and built environment features of cities can impact the access to resources that shape a healthy life such as good education, jobs, healthy diet, physical activity facilities, active transportation, etc. [25–27].

Loss of physical function, chronic pain, and the onset of chronic diseases result in worsening average SRH as populations age [22, 28]. Many studies analyzing the

association between aging and SRH have focused on elderly populations [23, 29–31]. However, the aging process is a continuum that can be explored throughout adulthood [32]. Characterizing SRH across the entire age span of adulthood may help us identify key modifiable conditions to improve the aging process.

Furthermore, several cross-sectional studies have shown that women tend to have poorer SRH than men [33–36]. Several mechanisms could explain these gender differences including gender-based roles in leisure and personal activities [33], higher prevalence of chronic diseases among women [36], differences in how men and women assess their health [37] and differential access to education and employment [34] as well as differential health returns to education in women and men [38, 39].

To our knowledge, very limited evidence is available regarding the association between age and SRH in Latin America cities, and the extent to which these relationships vary by gender or are modified by social and economic features of cities. An existing review of 11 studies in Latin America suggests that features of the built and social environment (where two studies used neighborhood socioeconomic variables) are associated with SRH or health-related quality of life [40]. However, the focus of these studies was mostly on smaller neighborhoods within cities, and most of the studies were small (samples ranging from 685 to 2045 subjects), and comprised few countries (only Brazil, Colombia and Cuba were represented in the 11 studies) [40]. Studies examining gender differences in the region have been also limited in scope (population > 60 years) [35]. Based on capabilities and economic welfare models, socioeconomic contexts of cities may impact SRH directly or may modify the aging effect on SRH (e.g. buffering or amplifying the impact) [41, 42]. However, the empirical evidence of these associations in cities in Latin America is lacking.

To fill these gaps, we examined the association between age and SRH across all cities of 100,000 residents or more in six Latin-American countries (Argentina, Brazil, Colombia, Chile, El Salvador, and Guatemala). We further evaluated if gender and city-level socioeconomic conditions are associated and/or modify the association between age and SRH. We used two main contextual social environment exposures: 1) a SALURBAL-derived Social Environment Index (SEI), which measures city-level material conditions of households (piped water, sewage system, overcrowding) and educational level which is a non-material household asset; and GDP per-capita, as a measure for cities SES, capturing dynamics of employment, growth, commerce, and business.

Methods

Study design and population

This is a cross-sectional, multi-level, multi-country study using data from the SALURBAL (Salud Urbana en América Latina) study. The SALURBAL study has integrated and harmonized health outcomes and physical and social environment data from 371 cities with more than 100,000 inhabitants in 11 Latin American countries [43]. The final analytical sample comprised 71,541 respondents, from 114 Latin-American cities in 6 countries (Additional file 1). The average number of observations per city was 628, ranging from 16 to 3487.

Health outcomes and socio demographic information at the individual level were obtained from existing national surveys in each country and harmonized across countries using standardized definitions. For the current analyses, we used data for adults (25–97 years) from the following countries surveys (year of survey): Argentina (2013), Brazil (2013), Chile (2010), Colombia (2007), El Salvador (2004), and Guatemala (2002) (Additional file 2), as these were the only SALURBAL harmonized survey data with SRH data. Colombia (2007) included adults ≤ 69 years. Social environment data — inputs for SEI— was derived from the following countries censuses (year of census): Argentina (2010), Brazil (2010), Chile (2002), Colombia (2005), El Salvador (2007), and Guatemala (2002) (Additional file 2).

The SALURBAL study protocol was approved by the Drexel University Institutional Review Board (IRB) with ID #1612005035 and by appropriate site-specific IRBs.

Outcome

SRH was measured on a 5-point Likert scale in each country. Although the question asked was fairly similar across countries, the response options differed. Adults were asked to answer the question “*In general, would you say your health is...*”. Argentina, Chile, Guatemala, and El Salvador used the exact same 5-point rating scale response including: “Poor”, “Fair”, “Good”, “Very Good” and “Excellent”. In Brazil and Colombia, the response rating scale included: “Very Poor”, “Poor”, “Fair”, “Good” and “Very Good”. During the data harmonization process, SRH was dichotomized into “Poor” (including “Very Poor”, “Poor”, “Fair”) and “Good” (including “Good”, “Very Good” and “Excellent”). This categorization facilitates comparison with other SRH epidemiological studies [19, 20, 29, 44–48].

Although SRH may be considered a more subjective measure of health compared to some objective measures of morbidity, we selected SRH as a proxy of global health status. SRH has shown to have high prognostic and predictive value of subsequent mortality and

therefore represents a good proxy of overall health [10, 11, 49]. SRH captures health status more holistically and is a potent proxy of peoples’ health related quality of life. Indeed, improving SRH constitutes an important and common goal of health systems, and urban health.

Exposures

Age (in years) was the primary individual-level exposure of interest and was obtained from the health surveys. Gender was our primary individual-level effect modifier, and it was inferred from self-reported sex from the survey responses (thus, we used sex as proxy for gender). We defined that if gender was an effect modifier, we would perform gender stratified analyses when investigating secondary effect modifiers.

The two co-primary city-level socioeconomic exposures included: Social Environment Index (SEI), and GDP per-capita.

The Social Environment Index, SEI, is a SALURBAL summary measure of city social development, which was already available at the time of the study [50]. SEI was created by combining four census-based indicators: education (% population with at least completed primary education among those aged 25 or above), water access (% households with access to piped water), sanitation (% households with access to a public sewage network) and reversed overcrowding (% households with more than 3 people per room). The index is the simple mean of these four city-level variables’ Z-scores. The four variables were selected because of their face validity in capturing different domains of the social environment and because of their predictive power of life expectancy differences across LATAM cities [50]. SEI has been used in previous research within the SALURBAL platform: for example, in relation to diabetes [51], hypertension [52] and green space [53]. Higher values of the city-SE index represent better city socioeconomic conditions.

GDP per-capita was obtained from a gridded global dataset [54], where subnational GDP estimates are based on an extensive econometric analysis [55] modeling GDP per capita for each city using a range of government, survey, and industry data. From this source annual gridded GDP (purchasing power parity in constant 2011 US dollars) per-capita, are available for the years from 1990 to 2015. To minimize the effect of annual fluctuations, we averaged 5 years, including the health survey year plus the 4 previous years.

Conceptually, these 2 city-level exposures can impact SRH directly or may modify the aging effect on SRH (e.g., buffering or amplifying the impact), thus we examined both the main effect of these exposures as well as their effect modification of the association between age and SRH.

Covariates

Individual education was included as proxy for individual socioeconomic level as a covariate. This was categorized as: “Less than Primary Education”, “Primary Education Completed”, “High-School Completed”, and “University Completed or a Higher level”. City size (number of people in the city) as estimated for the survey year based on country projections was also used in descriptive analyses.

Analysis

We used descriptive statistics to characterize the study population by outcome (Poor SRH and Good SRH), age categories (18–65 and 65+), SEI and GDP per-capita tertiles. We used continuous age (in years) divided by ten, so that coefficients reflect differences in SRH associated with 10 years of age. To accommodate the non-linear association between poor SRH and age, we used a linear spline with a knot at 65 years. We selected the knot at 65 years of age (coinciding with the retirement age in most countries), based on descriptive analyses of the relation between age and SRH. Furthermore, since the association differed by gender, all analyses were gender stratified.

We conducted two-level Poisson regression models, with individuals nested within cities, and with robust variance estimation to account for misspecification of the variance by fitting count model for the binary outcome [56]. We used Poisson models to estimate relative risk (RR) in place of logistic models since our outcome was common. We first examined effect modification by gender, and then fitted the following gender stratified models. Model 1: included age and country; Model 2: included model 1 plus individual-level education; Model 3: included model 2 plus GDP per-capita tertiles; Model 4: model 2 plus city-SEI tertiles and Model 5: model 2 plus both GDP per-capita and city-SEI tertiles. Country was added as fixed effect in all models to account for unmeasured features of countries that may confound the association of interest, and to account for differences in surveys years. To examine whether city-level factors modified the association between age and poor SRH, we further constructed a model which included the interactions of SEI and GDP per-capita (separately) with age and present the results graphically in both relative and absolute scales using linear combination of coefficients and adjusted marginal predicted prevalences, respectively. These models adjusted for individual-level education and had country as fixed effect.

Associations and interactions were considered significant if the p -values was <0.05 . The analyses were performed using STATA 16.1 (StataCorp, College Station, TX).

Results

The final analytical sample comprised 71,541 individuals, resident of 114 cities in 6 countries in Latin America. The mean age (SD) of the sample was 46.3 years (15.1) and 58% were women. Overall, 31% reported poor SRH. There were substantial variations in the proportion of poor SRH across the six countries. Those reporting poor SRH were older, more likely to be women and had lower educational attainment, as compared to those with good SRH. People reporting poor SRH were also different in terms of their city-level characteristics: they tended to live in cities with lower GDP per-capita, lower SEI, and smaller populations (Table 1).

Table 2 shows characteristics of the sample by age groups (18–65 and 65+). Compared to the younger group, older adults were more likely to be women (63% vs. 58%) and to report poor SRH (29% vs. 50%). Age distributions differed somewhat across countries. Argentina, Brazil, and Chile had a higher proportion in the older age group, while Colombia and Central America had larger proportion in the 18–65 age group. Older individuals had lower educational attainment but were more likely to live in cities with higher socioeconomic conditions as measured by SEI and GDP per-capita, and in cities with larger populations.

Geographical differences in city-level socioeconomic contexts are presented in Additional file 3. In general, Chile showed the best living conditions (piped water, sewage system and overcrowding), adults' education (completion of at least primary level) and GDP per-capita (mean and median). The lowest levels of SEI, GDP, education, and piped water are for Central America, who also bears the highest level of overcrowding (13.3%). Brazil is the country with the lower level of sewage system (58%) and the highest population (mean and median) per city.

Finally, the characteristics of the sample by city SEI and GDP per-capita tertile are shown in Additional file 4 and Additional file 5. Compared to tertiles 2 and 3 of SEI, individuals living in cities with lower social development (tertile 1) were younger, has a higher proportion of people aged 65+ and were less educated. They also had a higher proportion of people reporting poor SRH (38% in tertile 1 versus 29% in tertile 3). Also, cities in tertile 1 of SEI presented lower GDP per-capita and were smaller in terms of population (Additional file 4). Likewise, individuals living in cities with lower GDP per-capita (tertile 1) were slightly younger, had more people with less than primary education, and less people with university degree. They also had a higher proportion of people reporting poor SRH (36% versus 27% for tertile 3). Finally, cities in tertile 1 had the lower mean SEI and were of medium size population (Additional file 5).

Table 1 Characteristics of the study population by self-reported health status. SALURBAL Study ($n = 71,541$)

Characteristics	Poor SRH ($N = 22,480$)	Good SRH ($N = 49,061$)	Total ($N = 71,541$)	<i>p</i> -value*
Country contribution to sample				
% Argentina	19.6	28.8	26.0	< 0.001
% Brazil	48.8	44.4	45.8	
% Chile	4.6	2.8	3.4	
% Colombia	20.6	21.6	21.3	
% Guatemala & El Salvador	6.3	2.4	3.6	
Individual-level sociodemographic characteristics				
% Female	65.6	55.0	58.3	< 0.001
Mean (SD) Age in years	51.6 (15.5)	43.9 (14.2)	46.3 (15.1)	< 0.001
% 25–65 years	80.0	90.8	87.4	< 0.001
% > 65 years	20.0	9.2	12.6	
Educational level				
% Less than primary educ.	34.3	14.0	20.4	< 0.001
% Primary educ completed	33.9	27.6	29.6	
% High-School completed	24.8	38.8	34.4	
% University completed or higher level	7.0	19.5	15.6	
City-level socioeconomic characteristics				
Mean (SD) Socioeconomic Index, Z-score	−0.19 (0.97)	−0.02 (0.87)	0.17 (0.54)	< 0.001
Mean (SD) GDP per-capita	13,504.8 (8565.0)	14,769.6 (8937.9)	14,372.2 (8841.9)	< 0.001
Median GDP per-capita	10,401.6	11,225.4	11,225.4	
Mean (SD) Population size	3,549,314 (5027477)	3,754,636 (5330711)	3,690,119 (5238151)	
Median Population size	1,678,371	1,407,681	1,407,681	

GDP Gross Domestic Product, HH households, SD Standard deviation, SRH Self-rated health.

* *P* values from Wilcoxon Mann Whitney test for continuous variables and Fisher exact test for categorical variables

In an initial model pooling both genders and adjusting for education (not shown in table), age was non-linearly associated with poor SRH (risk ratio [RR]/10 years of age and 95% CI, 1.24 CI, 1.22,1.27 for people aged 25–65 and RR = 0.99 CI, 0.96, 1.0 for people aged 65 or more). Men had lower risk of poor SRH (RR = 0.76 CI 0.73–0.78), than women. Furthermore, gender modified the association of age with SRH (*p* values < 0.001 for persons < 65 and < 0.001 for persons 65 and over) such that men showed stronger increases associated with age than women. Therefore, all subsequent analyses were stratified by gender.

Tables 3 and 4 show associations of age, education, city GDP and city SEI with poor SRH in women ($n = 41,733$) and men ($n = 29,808$), respectively. There was a positive association between age and poor SRH in women and men, and in both genders the associations were stronger among people aged 25–65 (RR = 1.29 95% CI, 1.26,1.32; RR = 1.38 CI, 1.35,1.42, for women and men, respectively), than among those over 65 years (RR = 1.02 CI, 0.99,1.06; RR = 1.10 95 CI, 1.06,1.15, for women and men, respectively). After adjusting for individual education and city-level socio-economic

characteristics (models 2–5), age remained associated with poor SRH among those aged 25–65 years although the associations were slightly attenuated. In this age group, for each ten additional years the prevalence of reporting poor SRH increased 21% in women (RR = 1.21 CI, 1.18,1.24) and 30% in men (RR = 1.30 CI, 1.27,1.33). In contrast, age was not associated with poor SRH in women 65 years and above and was weakly associated with poor SRH in men 65 years or above (adjusted RR = 1.05 CI, 1.01,1.09). Model 3 for women and men shows that living in cities with less GDP per-capita (lowest tertile vs. the highest tertile) was associated with increased risk of poor SRH, even after adjusting for individual age and education level (RR = 1.24 CI 1.13,1.36; RR = 1.33 CI 1.18,1.49, for women and men respectively). Model 4 shows the adjusted association of tertiles of SEI and poor SRH, and demonstrates significant positive associations between cities with lower SEI (tertiles 1 and 2) and poor SRH compared to those with high SEI (tertile 3), in both men and women. To examine the independent association between SEI and GDP and SRH, we used a model that incorporates both factors (Model 5) and

Table 2 Characteristics of the study population by age groups. SALURBAL Study (N = 71,541)

Variables	Age group 1 N = 62,556	Age group 2 N = 8985	Difference p-value*
Age [range] in years	25–65	66–97	
Country contribution to sample			< 0.001
% Argentina	24.3	37.2	
% Brazil	45.6	47.1	
% Chile	3.1	5.2	
% Colombia	23.2	7.7	
% Guatemala & El Salvador	3.8	2.8	
Individual-level sociodemographic characteristics			
% Female	57.7	63.0	< 0.001
% Poor SRH	28.8	50.0	< 0.001
Educational attainment			
% Less than primary	16.9	44.7	< 0.001
% Primary Completed	29.7	28.9	
% High-School completed	37.0	16.2	
% University completed or higher level	16.4	10.1	
City-level socioeconomic characteristics			
Mean (SD) Socioeconomic Index, Z-score	−0.09 (0.92)	0.02 (0.78)	< 0.001
Mean (SD) GDP per-capita	14,080.12 (8782.2)	16,405.8 (8987.7)	< 0.001
Median GDP per-capita	11,225.4	16,263.2	
Mean (SD) Population size	3,601,395 (5139389)	4,307,837 (5843049)	< 0.001
Median Population size	1,407,681	1,646,057	

GDP Gross Domestic Product, SD Standard deviation, SRH Self-rated health.

* P values from Wilcoxon Mann Whitney test for continuous variables and Chi-square test for categorical variables

found that although there was some attenuation in the associations, these city-level variables remain significantly associated with poor SRH, for both men and women.

Results examining the role of SEI and GDP per-capita as effect modifiers are shown in Figs. 1, 2 respectively. Figure 1 shows adjusted marginal prevalences of SRH by age and SEI in each gender derived from gender specific models including interactions between age and SEI, and adjusting for individual education. We observed a more significant interaction for men in the younger group, with slightly stronger impact of increasing age on SRH among those living in cities with medium SEI (tertile 2).

Figure 2 for GDP shows, that among those aged 25–65 years, the association of age with poor SRH was stronger at higher than at lower levels of city GDP (p -value for interaction between age and GDP per-capita = 0.001 for women and < 0.001 for men). For men and women aged 65 and older no interaction between age and GDP was observed. Additional files 6 and 7, shows that results of the interaction model in the relative scale for SEI and GDP, respectively, are consistent with the results on the absolute scale.

Discussion

Based on cross-sectional health surveys from six Latin-American countries, we found that higher age was associated with poor SRH in men and women. In both genders the association between aging and poor SRH was stronger in middle-aged adults (25–65 years), than in older adults (65+) and remained significant even after accounting for individual level education and city-level socioeconomic characteristics. We further found that associations of age with poor SRH were stronger in men than in women. In both genders, both lower levels of city SEI and GDP per-capita were associated with poor SRH, even after adjusting for each other, with somewhat stronger associations for SEI as compared to GDP per capita. In addition, we found that GDP per-capita modified the associations between age and poor SRH, such that in middle-aged adults the age-SRH associations were stronger in cities with higher GDP per-capita. This interaction was less clear for city-SEI.

In our sample, after adjusting for education, the prevalence of poor SRH was 31.6% higher in women than in men. These findings are in line with previous studies showing that women are more likely to report poorer SRH levels than men [29–32]. The postulated

Table 3 Associations of age, education and city GDP with poor self-reported health in 114 cities in Latin America among women (n = 41,733)

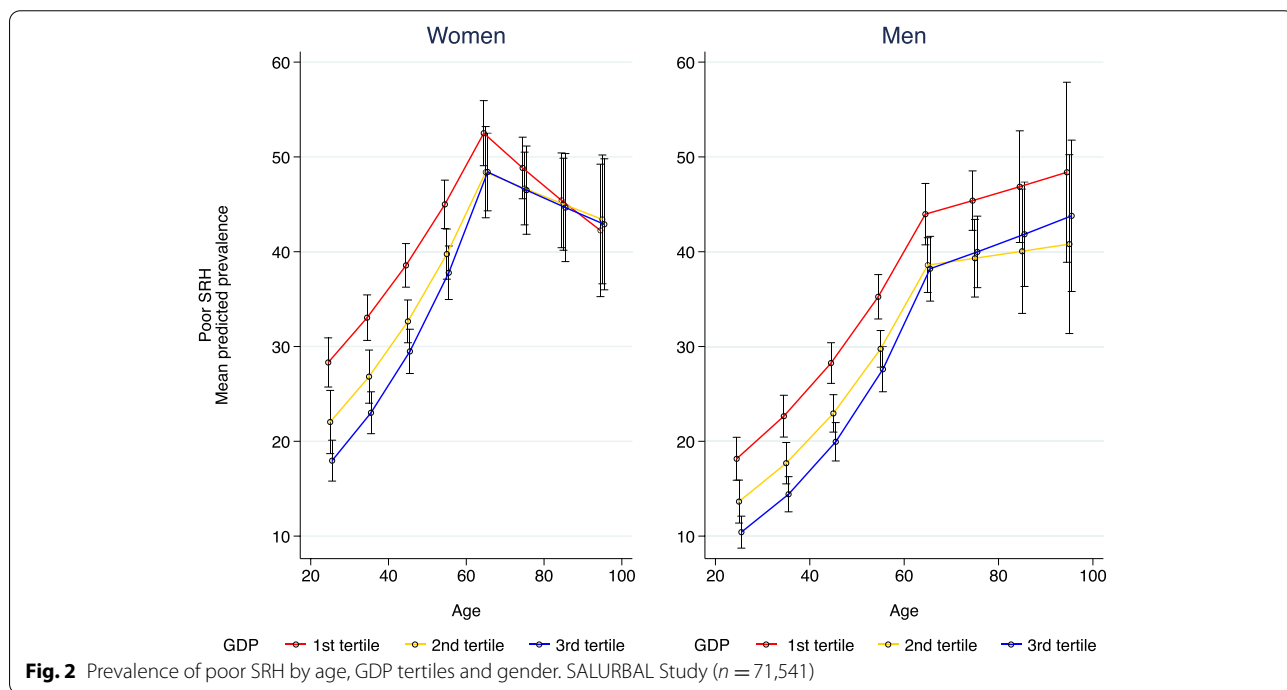
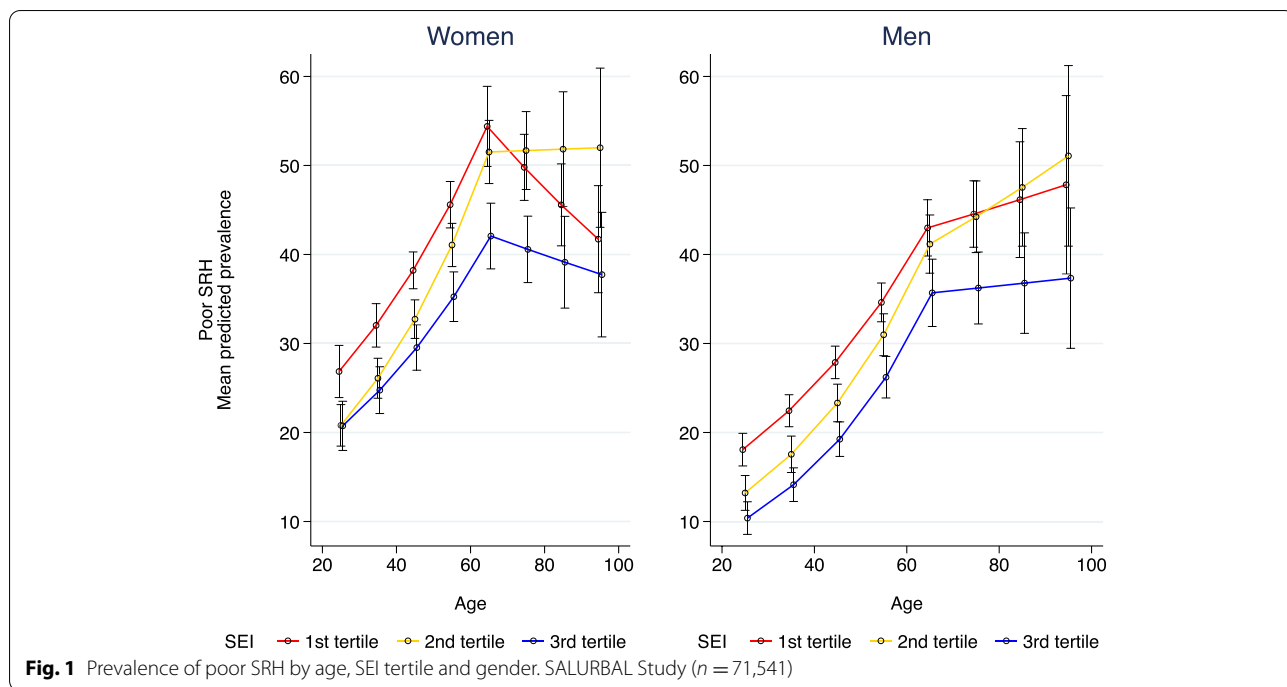
Variable	Model 1: Age		Model 2: Age, adjusting for individual education		Model 3: Model 2 + GDP Tertiles		Model 4: Model 2 + SEI Tertiles		Model 5: Model 2 + GDP tertiles + SEI Tertiles	
	RR (95% CI)	p-value	RR (95% CI)	p-value	RR (95% CI)	p-value	RR (95% CI)	p-value	RR (95% CI)	p-value
Individual-level factors										
Age, per 10years increase, among people aged 25–65 years	1.29 (1.26, 1.32)	<0.001	1.21 (1.18, 1.24)	<0.001	1.21 (1.18, 1.24)	<0.001	1.21 (1.18, 1.24)	<0.001	1.21 (1.18, 1.24)	<0.001
Age, per 10years increase, among people aged > 65 years	1.02 (0.99, 1.06)	0.19	0.97 (0.94, 1.00)	0.09	0.97 (0.94, 1.01)	0.09	0.97 (0.94, 1.01)	0.096	0.97 (0.94, 1.01)	0.096
Education less than primary			2.77 (2.59, 2.96)	<0.001	2.76 (2.58, 2.96)	<0.001	2.76 (2.58, 2.95)	<0.001	2.76 (2.58, 2.95)	<0.001
Education primary			2.39 (2.22, 2.57)	<0.001	2.39 (2.22, 2.57)	<0.001	2.39 (2.22, 2.57)	<0.001	2.39 (2.22, 2.57)	<0.001
Education secondary			1.60 (1.50, 1.71)	<0.001	1.60 (1.50, 1.71)	<0.001	1.60 (1.50, 1.71)	<0.001	1.60 (1.50, 1.71)	<0.001
Education university or higher			1.00 (reference)		1.00 (reference)		1.00 (reference)		1.00 (reference)	
City-GDP per capita and SEI										
GDP per capita: tertile 1					1.24 (1.13, 1.36)	<0.001			1.19 (1.13, 1.36)	0.005
GDP per capita: tertile 2					1.07 (0.98, 1.17)	0.15			1.06 (0.97, 1.16)	0.208
GDP per capita: tertile 3					1.00 (reference)				1.00 (reference)	
SEI: tertile 1									1.29 (1.17, 1.42)	<0.001
SEI: tertile 1									1.15 (1.04, 1.26)	0.004
SEI: tertile 1									1.00 (reference)	1.00 (reference)

All models consider country as fixed effect

Table 4 Associations of age, education, and city GDP with poor self-reported health in cities in 114 cities in Latin-America among men (n = 29,808)

Variable	Model 1: Age		Model 2: Age, adjusting for individual education		Model 3: Model 2 + GDP Tertiles		Model 4: Model 2 + SEI Tertiles		Model 5: Model 2 + GDP & SEI Tertiles	
	RR (95% CI)	p-value	RR (95% CI)	p-value	RR (95% CI)	p-value	RR (95% CI)	p-value	RR (95% CI)	p-value
Individual-level factors										
Age, per 10years increase, among people aged 25–65 years	1.38 (1.35, 1.42)	<0.001	1.30 (1.27, 1.33)	<0.001	1.30 (1.27, 1.33)	<0.001	1.30 (1.27, 1.33)	<0.001	1.30 (1.27, 1.33)	<0.001
Age, per 10years increase, among people aged > 65 years	1.10 (1.06, 1.15)	<0.001	1.05 (1.01, 1.09)	0.02	1.05 (1.01, 1.09)	0.02	1.05 (1.01, 1.09)	0.021	1.05 (1.01, 1.09)	0.022
Education: less than primary			3.20 (2.89, 3.55)	<0.001	3.20 (2.89, 3.55)	<0.001	3.19 (2.88, 3.55)	<0.001	3.19 (2.88, 3.55)	<0.001
Education: primary			2.64 (2.39, 2.92)	<0.001	2.64 (2.39, 2.92)	<0.001	2.64 (2.39, 2.91)	<0.001	2.64 (2.39, 2.91)	<0.001
Education: secondary			1.81 (1.64, 1.99)	<0.001	1.81 (1.64, 1.99)	<0.001	1.81 (1.64, 1.99)	<0.001	1.81 (1.64, 1.99)	<0.001
Education university or higher			1.00 (reference)		1.00 (reference)		1.00 (reference)		1.00 (reference)	
City-GDP per capita and SEI										
GDP per capita: tertile 1					1.33 (1.18, 1.49)	<0.001			1.27 (1.11, 1.46)	0.001
GDP per capita: tertile 2					1.09 (0.98, 1.22)	0.11			1.09 (0.98, 1.21)	0.117
GDP per capita: tertile 3					1.00 (reference)				1.00 (reference)	
SEI: tertile 1										
SEI: tertile 1									1.39 (1.24, 1.56)	<0.001
SEI: tertile 1									1.21 (1.07, 1.36)	0.002
SEI: tertile 1									1.00 (reference)	
									1.30 (1.17, 1.45)	<0.001
									1.24 (1.11, 1.39)	<0.001
									1.00 (reference)	

All models consider country as fixed effect



mechanisms behind these gender differences include differential exposure to factors that increase the risk of adverse health outcomes and differential vulnerability to these risk factors [57]. Some of the reasons for lower SRH in Latin America women could also lie in culture and

social structure. Women bear most of household chores, have fewer job opportunities and receive less income [35]. Women represent a high proportion of workers in the informal sector in Latin America. Informality, which exposes women to unsafe conditions including risk of

sexual harassment, often leaves women without the protection of labor laws, social benefits, health insurance or paid sick leave [58]. Further research is needed to better understand gender gaps in SRH in Latin America across adulthood.

Our results show that the impact of aging on SRH begins in early to mid-adulthood, we found stronger associations of higher age with poor SRH among young and middle-aged adults (25–65) than in older adults (65–97). Most previous studies analyzed the age effects on SRH focus on older population [22, 23, 29–31], and there are no comparable studies that explored age as main exposure for SRH across all adulthood. Additionally, cross-country comparisons in SRH cover mainly European or developed countries [12–14]. To our knowledge, there are two smaller studies, one from Giron, Spain and another one in Canada, with similar findings [45, 59]. The lack of association among older adults may be due to survival or selection bias, as people with worse health have higher mortality and those in poor health are unlikely to respond to surveys [22]. In our sample men 65+ showed a slight positive association of age with poor SRH, but this association was considerably less than that found in men aged 25–65 years. Birth cohort effects could also play a role in the associations of age with SRH that we report.

We also found that the association with age seemed to be stronger in men than in women. Reasons for these findings should be further studied in the region. A study from 2000, reported gender differences in SRH assessments in the elderly [37], noting that women tend to consider a broader range of information (beyond serious illnesses) when rating their health, while men would focus more on life threatening conditions [37], thus making SRH a strong mortality predictor for men only [37]. However, Zajacova et al. in a more recent (2017) and larger study for US individuals aged 25–84, concluded that SRH has similar meaning for men and women, and that both groups use a wide range of health-related information in forming their health judgements [60]. They found that women report worse SRH than men but only until mid-adulthood; the gender difference was reversed at older ages. They further reported that the excess of poor SRH among women disappears when differences in socio-economic and health covariates were considered [60]. Rohlfen & Kronenfeld, using longitudinal data for US population (mean age 56; SD 3 years), reported a significantly faster decline overtime in SRH in men as compared to women [57]. Transition to retirement, smoking status, and onset of chronic conditions could explain this faster decline of SRH in men. In women, lower baseline SRH was related to employment status and socioeconomic level; however, after adjustments for structural

and health status factors, women reported better SHR than men [57]. A qualitative study in 62 Wisconsin adults concluded that the way people take into account health issues to formulate their answers to the SRH question varies by socioeconomic status, gender and age [61]. An additional challenge in comparing age associations in men and in women using RR as we did is that the levels of poor SRH differ by gender being significantly higher in women than in men. Therefore, a smaller RR in women associated with age is compatible with a similar or even larger absolute difference (as suggested by Figs. 1 and 2). Further work is needed to better understand the differences in the associations of age with SRH in men and women and the extent to which they reflect differences in measurement, in exposures or in vulnerabilities by gender.

After adjusting for individual education, we found an independent association between both city-level SEI and GDP per-capita with SRH, in both men and women. These results suggest that these two social economic constructs have different mechanisms by which they impact SRH. SEI can exert a more direct influence of SRH as the index comprises crucial household assets (clean water, sewage system, dwelling space, and education), needed to accomplish basic levels of health and wellbeing, while GDP per-capita representing wealth, constitutes cities potential to achieve better health in a paradigm of economic welfare, perhaps, in a more indirect way. As stated in the introduction SEI approaches follow better (though not perfectly) the capabilities approach while GDP per-capita constitutes the classical welfarist approach to health and healthcare [41, 42]. One study at country level, found that the aggregated effect of education was twice the effect of income on cognitive functioning in adult aged 50 years and older, and thus indirectly moderated the effect of income on cognitive functioning. They further report that the effect sizes varied strongly between countries, concluding that country's GDP per-capita seems to influence cognitive functioning [62].

Other studies examining how urban contexts impact SRH have focused on neighborhood level factors [40, 63]. Within a review for Latin American countries, two out of 11 studies showed positive associations of neighborhoods socioeconomic levels and SRH, while others four showed that built environment features (parks, walkability, lack of noise) which are known to encourage healthy lifestyle, were associated to better SRH [40]. We did not explore built environment characteristics of the cities and therefore future research should widen the dimensions (and variables) in which the urban settings can affect and modify the age-SRH association in Latin America.

Finally, we found that GDP per-capita (more than SEI) modified the association of age and SRH in men and

women (25–65 years). Among people aged 25–65 we found stronger associations of age with SRH in cities with higher rather than lower GDP. The potential mechanisms for these results may include larger inequalities among cities with higher GDP per-capita and construct bias regarding self-rated health [12]. Which seems consistent with a recent study that used GDP per-capita at country level, reporting socioeconomic inequalities in physical and cognitive functions in a large sample (37 cohorts from 28 countries). The authors conclude that such inequalities exist across different social contexts, with varying magnitudes and appear to be larger in higher-income countries [64]. Further work is needed to understand how city GDP modifies exposures and vulnerabilities associated with aging.

Limitations and strengths

Our research has some limitations. Firstly, SRH may be considered a subjective measure of health compared to some objective measures of morbidity. Furthermore, to control for SES at the individual level, we relied only on educational level as a proxy for individual SES. Secondly, the cross-sectional nature of the study limits our ability to make causal inferences. Thirdly, although we had few missing data (4.3% - see Additional file 1), country responders may not be representative of all health status and age strata within included cities. The potential underrepresentation of elderly people, and people with severe conditions or multimorbidity might introduce some bias towards good SRH. Moreover, one of the country surveys (Colombia) only included people under 70 years of age.

We also encountered lack of alignment in the calendar-years of the data sources, e.g. year of the survey (source of SRH and individual covariates) and the census year (source of the SEI). The gap between survey data and census was largest for Chile (8-year difference), and smallest for Colombia, Argentina, Brazil and El Salvador (2–3 year difference). Although there may have been differences in the prevalence of poor SRH over time, we do not anticipate differences in the associations that were examined in this study. Another related limitation is the fact that Guatemala data is the oldest (2002), but similarly, we do not expect changes in the associations over time. Finally, the role of built environment characteristics of the cities (such as air pollution, greenness, climate, or traffic) on SRH, was beyond the current scope of this manuscript. Future studies in SALURBAL are very well positioned to address this research question.

Despite these shortcomings, this study has important strengths. This is one of the largest studies to date, with a harmonized dataset, focusing on a wide and heterogeneous range of Latin American cities (114 cities from 6 countries), most of the countries are classified in the

category of middle-income by the World Bank. Our sample includes adults with a wide range of ages and thus extends our knowledge of SRH across adulthood. Finally, the study multi-level structure, nesting people within cities, allows accounting for unmeasured city contextual factors which may have a role in shaping SRH as people age.

Conclusions

In Latin American cities aging was associated with higher risk of poor SRH, but this association was stronger in adults from 25 to 65 than in those over 65. Although women are more likely to report poor SRH, we found that the association between aging and poor SRH was more pronounced among men. City SEI and GDP per-capita were associated with SRH with stronger and most consistent effects for SEI. GDP per-capita modified the association between age and poor SRH such that high GDP was associated with stronger age association.

The main policy and research implications are the need to better understand the gender differences in the experience of perceived health with aging in Latin American cities. Women rate their health worse than men, but men feel a stronger deterioration in SRH as they age. Future research could explore the extent to which gender-based cultural, social, or political issues shape these trajectories in Latin American cities. In addition, more research (qualitative and quantitative) is needed to better understand how city socioeconomic environments represented by different indicators (such material resources vs. GDP per-capita) exert potentially independent influences on the age effect on SRH, suggesting different mechanisms behind these associations. Studies should also investigate what Latin American people consider when rating their health. On the other hand, realizing that the age effect on SRH is stronger in younger adults rather than people over 65 years is a key finding that should guide the design of better policies that support healthy adulthood in urban settings.

Additionally, actions to improve (perceived) health status should begin at early ages in adulthood considering the different gender trajectories as people age and the role of cities in delaying and attenuating the impact of age in SRH in Latin American countries.

Abbreviations

GDP per-capita: Gross Domestic Product per-capita; Good SRH: good or better self-reported or self-rated health; HRQoL: Health-related Quality-of-Life; IRB: Institutional Review Board; LATAM: Latin American; Poor SRH: Poor self-rated health (includes the categories of fair, poor and very poor self-reported or self-rated health); SALURBAL: Salud Urbana en América Latina; SEI: Socioeconomic Index; SRH: Self-reported or self-rated health.

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12889-022-13752-2>.

Additional file 1. Derivation of the analytical sample. SALURBAL Study

Additional file 2. Data sources per country and analytical sample (individuals and cities)

Additional file 3. Socioeconomic characteristics of the cities in the study, by country

Additional file 4. Characteristics of the study population by SEI tertiles. SALURBAL Study ($N = 71,541$)

Additional file 5. Characteristics of the study population by GDP per capita (in constant 2011 USD) tertiles. SALURBAL Study ($N = 71,541$)

Additional file 6. Association between individual age and poor self-reported health by tertiles of Socioeconomic Index, age, and gender.

Additional file 7. Association between individual age and poor self-reported health by tertiles of GDP, age, and gender

Acknowledgements

The authors acknowledge the contribution of all SALURBAL project team members. For more information on SALURBAL see <https://drexel.edu/lac/salurbal/team/>.

SALURBAL project team members:

Marcio Alazraqui, Hugo Spinelli, Carlos Guevel, Vanessa Di Cecco, Adela Tisn s, Carlos Leveau, Adrian Santoro, Damian Herkovits, Andres Trotta, Patricia Aguirre, Serena onica Perner: National University of Lanus, Buenos Aires, Argentina; Santiago Rodriguez Lopez, Natalia Tumas: CIECS - Centro de Investigaciones y Estudios sobre Cultura y Sociedad, Universidad Nacional de Córdoba, Córdoba, Argentina; Nelson Gouveia, Maria Antonietta Mascoll, Anne Dorothe Slovic, Lucas Soriano Martins, Claudio Makoto Kanai: Universidade de Sao Paulo, Sao Paulo, Brazil; Mauricio Barreto, Gervasio Santos, Anderson Dias de Freitas, Aureliano Sancho Souza Paiva, Jose Firmino de Sousa Filho, Maria Izabel dos Santos Bell, Roberto Fernandes Silva Andrade: Oswaldo Cruz Foundation, Salvador Bahia, Brazil; Caio Porto De Castro: Universidade Federal da Bahia, Salvador Bahia, Brazil; Letícia de Oliveira Cardoso, Mariana Carvalho de Menezes, Maria de Fatima Rodrigues Pereira de Pina, Daniel Albert Skaba, Joanna Miguez Nery Guimaraes, Vanderlei Pascoal de Matos: Oswaldo Cruz Foundation, Rio de Janeiro, Brazil; Mariana Carvalho de Menezes: Universidade Federal de Ouro Preto, Ouro Preto, Brazil; Waleska Teixeira Caiaffa, Amélia Augusta de Lima Friche, Carina Maris de Souza, Dora Moraes Coelho, Denise Marques Sales, Guilherme Aparecido Santos Aguiar, Guilherme Ottoni, Julia de Carvalho Nascimento, Lidia Maria de Oliveira Moraes, Mariana de Melo Santos, Solimar Carnavalli Rocha, Uriel Moreira Silva: Universidade Federal de Minas Gerais, Belo Horizonte, Brazil; Camila Teixeira Vaz: Universidade Federal de Juiz de Fora, Juiz de Fora, Brazil; Amanda Cristina de Souza Andrade: Universidade Federal de Mato Grosso, Mato Grosso, Brazil; Patricia Frenz, Tania Alfaro, Carolina Nazzal, Cynthia Córdoba, Pablo Ruiz, Mauricio Fuentes, Marianela Castillo, Rodrigo Mora, Sebastian Pedrero, Lorena Rodriguez, Sandra Flores, Tamara Doberti: School of Public Health, University of Chile, Santiago, Chile; Alejandra Vives Vergara, Alejandro Salazar, Cristina Schmitt, Daniela Olivares, Francisca Gonzalez, Fernando Baeza, Flavia Angelini, Ignacio Diaz, Laura Orlando, Natalia Diaz, Pablo Campos, Roxana Valdebenito, Victoria Leon: Department of Public Health, Pontificia Universidad Catolica de Chile, Santiago de Chile; Andrea Cortinez-O'Ryan: Universidad de La Frontera, Santiago de Chile; Olga Lucia Sarmiento, Andres Felipe Aguilar, Julian Arellana, Claudia Bedoya, Jorge Alexander Bonilla, Marcelo Botero, Sergio Cabrales, German Carvajal, Natalia Cely, Diego Lucumi Cuesta, Carlos Mauricio Diaz, Karen Fajardo, Catalina Gonzalez, Silvia Alejandra Gonzalez, Oscar Guaje, John Alexis Guerra, Paula Guevara, Tomas Guevara Aladino, Luis Angel Guzman, Philipp Hessel, Diana Higuera, Bernardo Huertas, Jorge Huertas, Ana Maria Jaramillo, Joaquin Hernando Jaramillo Sabogal, Mario Linares, Julieth Lopez, Diego Lucumi, Paola Martinez, Andres Medaglia, Daniela Mendez, Ricardo Morales, Felipe Montes, Anamaria Muñoz Florez, Alejandro Palacio, Fabian Camilo Peña, Jose David Pinzon, Camilo Triana, Andres Felipe Useche, Maria Alejandra Wilches, Sandra Zuñiga: Universidad de los Andes, Bogota, Colombia; Carlos Moncada: Universidad Nacional, Bogota, Colombia; Lina Martinez: Universidad Icesi, Cali, Colombia;

Jose David Meisel: Universidad de Ibagu, Ibagu, Colombia; Eliana Martinez: Universidad de Antioquia, Medellin, Colombia; Maria Fernanda Kroker-Lobos, Manuel Ramirez-Zea, Monica Mazariegos, Anali Morales: INCAP Research Center for the Prevention of Chronic Diseases (CIPEC), Institute of Nutrition of Central America and Panama (INCAP), Guatemala City, Guatemala; Tonatiuh Barrientos-Gutierrez, Arantxa Colchero Aragones, Carolina Perez-Ferrer, Francisco Javier Prado-Galbarro, Nancy Paulina Lopez Olmedo, Filipa de Castro, Rosalba Rojas-Martinez, Alejandra Jauregui, Dalia Stern, Horacio Riojas, Jose Luis Texcalac, Herney Alonso Rengifo Reina, Desiree Vidana Perez, Yenisei Ramirez Toscano: Instituto Nacional de Salud Publica, Mexico City, Mexico; J. Jaime Miranda, Cecilia Anza-Ramirez, Francisco Diez-Canseco, Akram Hernandez Vasquez, Lorena Saavedra-Garcia, Jessica H. Zafra-Tanaka (CRONICAS Centre of Excellence in Chronic Diseases, Universidad Peruana Cayetano Heredia, Lima, Peru); Ross Hammond: Brookings Institute, Washington, D.C., USA; Daniel Rodriguez, Maryia Bakhtsiyarava, Iryna Dronova, Xize Wang, Mika Ruchama Moran, Yuanyuan Zhao, Yang Ju, Xavier Delclos-Ali: Department of City and Regional Planning, the University of California at Berkeley, California, USA; Peter Hovmand, Ellis Ballard, Jill Kuhlberg: Washington University in St Louis, St. Louis, Missouri, USA; Ana Diez Roux, Binod Acharya, Amy Auchincloss, Ione Avila-Palencia, Sharrelle Barber, Usama Bilal, Ariela Braverman, Dustin Fry, Felipe Garcia-España, Katherine Indvik, Josiah Kephart, Carolyn Knoll, Brent Langelier, Mariana Lazo, Ran Li, Gina Lovasi, Rosie Mae Henson, Kevin Martinez-Folgar, Steve Melly, Yvonne Michael, Kari Moore, Jeff Moore, Prilica Mullachery, Ana Ortigoza, Harrison Quick, D. Alexander Quistberg, Jordan Rodriguez Hernandez, Brisa Sanchez, S. Claire Slesinski, Ivana Stankov, Jose Tapia Granados, Bricia Trejo, Goro Yamada: Urban Health Collaborative, Dornsife School of Public Health, Drexel University, Philadelphia, Pennsylvania, USA.

SALURBAL acknowledges the contributions of many different agencies in generating, processing, facilitating access to data or assisting with other aspects of the project. Please visit <https://drexel.edu/lac/data-evidence/data/> for a complete list of data sources. The findings of this study and their interpretation are the responsibility of the authors and do not represent the views or interpretations of the institutions or groups that compiled, collected, or provided the data. The use of data from these institutions does not claim or imply that they have participated in, approved, endorsed, or otherwise supported the development of this publication. They are not liable for any errors, omissions or other defect or for any actions taken in reliance thereon.

Authors' contributions

MCR helped conceive the research, conducted the analyses, drafted and revised the article. GY and ML helped conceive the study, supervised the statistical analysis, assisted with drafting and critically reviewed several versions of the article. AVDR secured funding and helped conceive for the study, assisted with drafting and critically reviewed the different versions of the article. TA and SFA helped conceive, assisted with the analyses and critically reviewed several versions of the article. TB secured funding for the study, helped conceive the research, assisted with the initial analyses and critically reviewed the article. OLS secured funding for the study, helped conceive the research and critically reviewed the article. CTV helped conceive the research and critically reviewed the article. All authors read and approved the final manuscript.

Funding

The Salud Urbana en América Latina (SALURBAL) / Urban Health in Latin America project is funded by the Wellcome Trust [205177/Z/16/Z]. The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript. Marianela Castillo-Riquelme is supported by the Chilean National Research and Development Agency through a PhD scholarship program (Beca Doctorado Nacional 2019/21190128) and by the SALURBAL Project as part-time researcher at the University of Chile.

Availability of data and materials

The SALURBAL project welcomes queries from anyone interested in learning more about its dataset and potential access to data. To learn more about SALURBAL visit <https://Drexel.edu/lac/> or contact the project at salurbal@drexel.edu.

Declarations

Ethics approval and consent to participate

The SALURBAL study protocol was approved by the Drexel University Institutional Review Board with ID #1612005035 and by appropriate site-specific IRBs.

All methods were performed in accordance with the relevant guidelines and regulations. For health survey participants, informed consent was obtained from all subjects.

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

Author details

¹Doctoral Program in Public Health, School of Public Health, Faculty of Medicine, University of Chile, Avenida Independencia, 939 Santiago, Chile. ²Urban Health Collaborative, Dornsife School of Public Health, Drexel University, Philadelphia, PA, USA. ³Instituto Nacional de Salud Pública, Cuernavaca, Mexico. ⁴Campus Centro-Oeste Dona Lindu, Federal University of São João del-Rei, Divinópolis, Brazil. ⁵Institute of Collective Health, National University of Lanus, Buenos Aires, Argentina. ⁶School of Medicine, Universidad de los Andes, Bogota, Colombia. ⁷Department of Community Health and Prevention, Dornsife School of Public Health, Drexel University, Philadelphia, PA, USA.

Received: 2 March 2022 Accepted: 4 July 2022

Published online: 05 August 2022

References

- United nations. World Urbanization Prospects 2018 Highlights 2019.
- United Nations, Department of Economic and Social Affairs PD. World Urbanization Prospects: The 2018 Revision, Online Edition (File 21). [Internet]. 2018 [cited 2022 May 10]. Available from: <https://population.un.org/wup/Download/>
- United Nations Department of Economic and Social Affairs Population Division. World Population Ageing 2019. Highlights. Vol. Highlights. NY: World Population Ageing 2019; 2019. p. 64. <https://www.un.org/en/development/desa/population/publications/pdf/ageing/WorldPopulationAgeing2019-Report.pdf>.
- Sudharsanan N, Bloom DE. The Demography of Aging in Low- and Middle-Income Countries: Chronological versus Functional Perspectives. In: National Academies of Sciences, Engineering and MD of B and SS and EC on P. Washington (DC): Future Directions for the Demography of Aging: Proceedings of a Workshop. National A; 2018.
- WHO. Global age-friendly cities: A guide, vol. 87. Bristol: World Health Organization; 2007.
- OECD. Ageing in Cities. Paris: OECD Publishing; 2015. <https://doi.org/10.1787/9789264231160-en>.
- Cleary PD. Subjective and objective measures of health: which is better when? J Heal Serv Res Policy. 1997;2(1):3–4.
- Lundberg O, Manderbacka K. Assessing reliability of a measure of self-rated health. Scand J Soc Med. 1996;24(3):218–24.
- Kaplan SH. Patient reports of health status as predictors of physiologic health measures in chronic disease. J Chronic Dis. 1987;40:275–355.
- Mossey JM, Shapiro E. Self-rated health: a predictor of mortality among the elderly. Am J Public Health. 1982;72(8):800–8.
- Mavaddat N, Valderas JM, Van Der Linde R, Khaw KT, Kinmonth AL. Association of self-rated health with multimorbidity, chronic disease and psychosocial factors in a large middle-aged and older cohort from general practice: A cross-sectional study. BMC Fam Pract. 2014;15(1):185.
- Jürges H. True health vs response styles: exploring cross-country differences in self-reported health. Health Econ. 2007;16(2):163–78.
- Gray L, Merlo J, Mindell J, Hallqvist J, Tafforeau J, O'Reilly D, et al. International differences in self-reported health measures in 33 major metropolitan areas in Europe. Eur J Pub Health. 2012;22(1):40–7.
- Balaj M, McNamara CL, Eikemo TA, Bamba C. The social determinants of inequalities in self-reported health in Europe: findings from the European social survey (2014) special module on the social determinants of health. Eur J Public Health. 2017;27(suppl_1):107–14.
- Wang L, Dong W, Ou Y, Chen S, Chen J, Jiang Q. Regional differences and determinants of self-rated health in a lower middle income rural Society of China. Int J Equity Health. 2018;17(1):162.
- Fernandez-Martinez B, Prieto-Flores M-E, Forjaz MJ, Fernández-Mayoralas G, Rojo-Pérez F, Martínez-Martín P. Self-perceived health status in older adults: regional and sociodemographic inequalities in Spain. Rev Saude Publica. 2012;46(2):310–9.
- Subramanian V, Delgado I, Jadue L, Kawachi I, Vega J. Income inequality and self rated health: an analysis from a contextual perspective in Chile. Rev Med Chil. 2003;131(3):321–30.
- Caicedo B, Berbesi FD. Self-rated health in adults: influence of poverty and income inequality in the area of residence. Gac Sanit. 2015;29(2):97–104.
- Subramanian SV, Kawachi I. Whose health is affected by income inequality? A multilevel interaction analysis of contemporaneous and lagged effects of state income inequality on individual self-rated health in the United States. Heal Place. 2006;12(2):141–56.
- Dunstan F, Fone DL, Glickman M, Palmer S. Objectively measured residential environment and self-reported health: A multilevel analysis of UK census data. PLoS One. 2013;8(7):e69045.
- Oshio T, Kobayashi M. Income inequality, area-level poverty, perceived aversion to inequality, and self-rated health in Japan. Soc Sci Med. 2009;69(3):317–26.
- Andersen FK, Christensen K, Frederiksen H. Self-rated health and age: A cross-sectional and longitudinal study of 11,000 Danes aged 45–102. Scand J Public Health. 2007;35(2):164–71.
- Liang J, Shaw BA, Krause N, Bennett JM, Kobayashi E, Fukaya T, et al. How does self-assessed health change with age? A study of older adults in Japan. J Gerontol B Psychol Sci Soc Sci. 2005;60(4):S224–32.
- Volken T, Wieber F, Ruesch P, Huber M, Crawford RJ. Temporal change to self-rated health in the Swiss population from 1997 to 2012: the roles of age, gender, and education. Public Health. 2017;150:152–65.
- Yeager RA, Smith TR, Bhatnagar A. Green environments and cardiovascular health. Trends Cardiovasc Med. Elsevier Inc. 2020;30(4):241–6.
- Orr MG, Kaplan GA, Galea S. Neighbourhood food, physical activity, and educational environments and black/white disparities in obesity: a complex systems simulation analysis. J Epidemiol Community Health. 2016;70(9):862–7.
- Passi-Solar Á, Margozzini P, Cortinez-O'Ryan A, Muñoz JC, Mindell JS. Nutritional and metabolic benefits associated with active and public transport: results from the Chilean National Health Survey, ENS 2016–2017. J Transp Heal. 2020;17:100819.
- Hoeymans N, Feskens EJM, Van Den Bos GAM, Kromhout D. Age, time, and cohort effects on functional status and self-rated health in elderly men. Am J Public Health. 1997;87(10):1620–5.
- Girón P. Is age associated with self-rated health among older people in Spain? Cent Eur J Public Health. 2012;20(3):185–90.
- Ocampo-Chaparro JM, Zapata-Ossa HD, Cubides-Munévar AM, Curcio CL, Villegas JD, Reyes-Ortiz CA. Prevalence of poor self-rated health and associated risk factors among older adults in Cali, Colombia. Colomb medica (Cali). 2013;44(4):224–231.
- Spuling SM, Wurm S, Tesch-Römer C, Huxhold O. Changing predictors of self-rated health: disentangling age and cohort effects. Psychol Aging. 2015;30(2):462–74.
- Jamoom EW, Horner-Johnson W, Suzuki R, Andresen EM, Campbell VA. Age at disability onset and self-reported health status. BMC Public Health. 2008;8(1):10.
- Adjei NK, Brand T, Zeeb H. Gender inequality in self-reported health among the elderly in contemporary welfare countries: A cross-country analysis of time use activities, socioeconomic positions and family characteristics. Assari S, editor. PLoS One. 2017;12(9):e0184676.
- Boerma T, Hosseinpoor AR, Verdes E, Chatterji S. A global assessment of the gender gap in self-reported health with survey data from 59 countries. BMC Public Health. 2016;16:675.
- Trujillo AJ, Mroz TA, Piras C, Vernon JA, Angeles G. Determinants of gender differences in health among the elderly in Latin America. World Health Popul. 2010;11(3):24–43.

36. Malmusi D, Artazcoz L, Benach J, Borrell C. Perception or real illness? How chronic conditions contribute to gender inequalities in self-rated health. *Eur J Pub Health*. 2012;22(6):781–6.
37. Benyamini Y, Leventhal EA, Leventhal H. Gender differences in processing information for making self-assessments of health. *Psychosom Med*. 2000;62(3):354–64.
38. Ross CE, Masters RK, Hummer RA. Education and the gender gaps in health and mortality. *Demography*. 2012;49(4):1157–83.
39. Zhu B, Ye Y. Gender disparities in the education gradient in self-reported health across birth cohorts in China. *BMC Public Health*. 2020;20(1):1–11.
40. Gomez LF, Soto-Salazar C, Guerrero J, Garcia M, Parra DC. Neighborhood environment, self-rated health and quality of life in Latin America. *Health Promot Int*. 2020;35(2):196–204.
41. Wells T. Sen's Capability Approach [Internet]. *Internet Encyclopedia of Philosophy*. 2022 [cited 2022 May 10]. Available from: <https://iep.utm.edu/sen-cap/>
42. Hall PA, Taylor RCR, Barnes L. A capabilities approach to population health and public policy-making. *Rev Epidemiol Sante Publique*. 2013;61(SUPPL. 3):S177–83.
43. Quistberg DA, Diez Roux AV, Bilal U, Moore K, Ortigoza A, Rodriguez DA, et al. Building a data platform for cross-country urban health studies: the SALURBAL study. *J Urban Heal*. 2019;96(2):311–37.
44. Manor O, Matthews S, Power C. Dichotomous or categorical response? Analysing self-rated health and lifetime social class. *Int J Epidemiol*. 2000;29(1):149–57.
45. Shooshtari S, Menec V, Tate R. Comparing predictors of positive and negative self-rated health between younger (25–54) and older (55+) Canadian adults: A longitudinal study of well-being. *Res Aging*. 2007;29(6):512–54.
46. Han S, Kim H, Lee H-S. A multilevel analysis of social capital and self-reported health: evidence from Seoul, South Korea. *Int J Equity Health*. 2012;11(1):3.
47. Chaparro MP, Hughes A, Kumari M, Benzeval M. Is the association between self-rated health and underlying biomarker levels modified by age, gender, and household income? Evidence from understanding society – the UK household longitudinal study. *SSM - Popul Heal*. 2019;8:100406.
48. Witvliet MI, Kunst AE, Stronks K, Arah OA. Variations between world regions in individual health: A multilevel analysis of the role of socio-economic factors. *Eur J Pub Health*. 2012;22(2):284–9.
49. DeSalvo KB, Blosner N, Reynolds K, He J, Muntner P. Mortality prediction with a single general self-rated health question. A meta-analysis. *J Gen Intern Med*. 2006;21(3):267–75.
50. Bilal U, Hessel P, Perez-Ferrer C, Michael YL, Alfaro T, Tenorio-Mucha J, et al. Life expectancy and mortality in 363 cities of Latin America. *Nat Med*. 2021;27(3):463–70.
51. Braverman-Bronstein A, Hessel P, González-Uribe C, Kroker MF, Diez-Canseco F, Langellier B, et al. Association of education level with diabetes prevalence in Latin American cities and its modification by city social environment. *J Epidemiol Community Health*. 2021;75(9):874–80.
52. Avila-Palencia I, Rodríguez DA, Miranda JJ, Moore K, Gouveia N, Moran MR, et al. Associations of urban environment features with hypertension and blood pressure across 230 Latin American cities. *Environ Health Perspect*. 2022;130(2):1–10.
53. Ju Y, Moran M, Wang X, Avila-Palencia I, Cortinez-O'Ryan A, Moore K, et al. Latin American cities with higher socioeconomic status are greening from a lower baseline: evidence from the SALURBAL project. *Environ Res Lett*. 2021;16(10):104052.
54. Kummu M, Taka M, Guillaume JHA. Gridded global datasets for gross domestic product and human development index over 1990–2015. *Sci Data*. 2018;5:1–15.
55. Gennaioli N, LA-Porta R, Lopez-de-Silanes F, Shleifer A. Human capital and regional development. *Q J Econ*. 2013;128(1):105–64.
56. Zou G. A modified Poisson regression approach to prospective studies with binary data. *Am J Epidemiol*. 2004;159(7):702–6.
57. Rohlfen LS, Jacobs KJ. Gender differences in trajectories of self-rated health in middle and old age: an examination of differential exposure and differential vulnerability. *J Aging Health*. 2014;26(4):637–62.
58. UN Women. *Progress of the World's Women 2015–2016*. Chapter 2. p. 71. Vol. 148.
59. Girón P. Determinants of self-rated health in Spain: differences by age groups for adults. *Eur J Pub Health*. 2012;22(1):36–40.
60. Zajacova A, Huzurbazar S, Todd M. Gender and the structure of self-rated health across the adult life span. *Soc Sci Med*. 2017;187:58–66.
61. Garbarski D, Dykema J, Croes KD, Edwards DF. How participants report their health status: cognitive interviews of self-rated health across race/ethnicity, gender, age, and educational attainment. *BMC Public Health*. 2017;17(1):1–13.
62. Rodriguez FS, Hofbauer LM, Röhr S. The role of education and income for cognitive functioning in old age: A cross-country comparison. *Int J Geriatr Psychiatry*. 2021;36(12):1908–21.
63. Behanova M, Reijneveld SA, Nagyoova I, Katreniakova Z, van Ameijden EJC, Dijkshoorn H, et al. Are area-level and individual-level socioeconomic factors associated with self-rated health in adult urban citizens? Evidence from Slovak and Dutch cities. *Eur J Pub Health*. 2017;27(suppl_2):86–92.
64. Stefler D, Prina M, Wu YT, Sánchez-Niubò A, Lu W, Haro JM, et al. Socioeconomic inequalities in physical and cognitive functioning: cross-sectional evidence from 37 cohorts across 28 countries in the ATHLOS project. *J Epidemiol Community Health*. 2021;75(10):980–6.

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Ready to submit your research? Choose BMC and benefit from:

- fast, convenient online submission
- thorough peer review by experienced researchers in your field
- rapid publication on acceptance
- support for research data, including large and complex data types
- gold Open Access which fosters wider collaboration and increased citations
- maximum visibility for your research: over 100M website views per year

At BMC, research is always in progress.

Learn more biomedcentral.com/submissions

